Statistical Theory;
Why is the Gaussian Distribution so popular?

Rob Nicholls

MRC LMB Statistics Course 2014
Contents

• Continuous Random Variables
• Expectation and Variance
• Moments
• The Law of Large Numbers (LLN)
• The Central Limit Theorem (CLT)
A *Random Variable* is an object whose value is determined by chance, i.e. random events.

Probability that the random variable $X$ adopts a particular value $x$:

$$P(X = x) = \begin{cases} 
  p \in [0,1] & X: \text{discrete} \\
  0 & X: \text{continuous}
\end{cases}$$
Continuous Uniform Distribution: \( X \sim U(a,b) \)

Probability Density Function:
\[
f_X(x) = \begin{cases} 
  (b-a)^{-1} & \text{if } x \in [a,b] \\
  0 & \text{if } x \notin [a,b] 
\end{cases}
\]
Continuous Random Variables

Example: \( X \sim U(0,1) \)

\[
P(X \in [0,1]) = 1
\]
Continuous Random Variables

Example: \( X \sim U(0,1) \)

\[
P(X \in [0, \frac{1}{2}]) = \frac{1}{2}
\]
Continuous Random Variables

Example: \( X \sim U(0,1) \)

\[
P(X \in [0, \frac{1}{3}]) = \frac{1}{3}
\]
Continuous Random Variables

Example: \( X \sim U(0,1) \)

\[ P(X \in [0, \frac{1}{10}]) = \frac{1}{10} \]
Continuous Random Variables

Example:  \( X \sim U(0,1) \)

\[
P(X \in [0, \frac{1}{100}]) = \frac{1}{100}
\]
Continuous Random Variables

Example: \( X \sim U(0,1) \)

\[
P(X \in [0, \frac{1}{n}]) = \frac{1}{n}
\]

\[
\lim_{n \to \infty} P(X \in [0, \frac{1}{n}]) = 0
\]

\[
\lim_{n \to \infty} P(0 \leq X \leq \frac{1}{n}) = 0
\]

In general, for any continuous random variable \( X \):

\[
\lim_{\varepsilon \to 0} P(0 \leq X \leq \varepsilon) = 0
\]

\[
\lim_{\varepsilon \to 0} P(\alpha \leq X \leq \alpha + \varepsilon) = 0
\]
Continuous Random Variables

\[ P(X = x) = \begin{cases} 
  p_X(x) & \text{if } X : \text{discrete} \\
  0 & \text{if } X : \text{continuous} 
\end{cases} \]

“Why do I observe a value if there’s no probability of observing it?!”

Answers:
- Data are discrete
- You don’t actually observe the value – precision error
- Some value must occur... even though the probability of observing any particular value is infinitely small
Continuous Random Variables

For a random variable: \( X : \Omega \rightarrow A \)

The *Cumulative Distribution Function (CDF)* is defined as:

\[
F_X(x) = P(X \leq x) \quad \text{(discrete/continuous)}
\]

Properties:

- Non-decreasing
- \( \lim_{x \rightarrow -\infty} F_X(x) = 0 \)
- \( \lim_{x \rightarrow +\infty} F_X(x) = 1 \)
Continuous Random Variables

Probability Density function:

Cumulative Distribution function:

Boxplot:
Continuous Random Variables

Probability Density function:

$$f_X(x)$$

Cumulative Distribution function:

$$F_X(x) = P(X \leq x)$$

Boxplot:
Continuous Random Variables

Probability Density function:

$\int_{-\infty}^{x} f_X(y) dy$

Cumulative Distribution function:

$F_X(x) = P(X \leq x)$

Boxplot:
Continuous Random Variables

Probability Density function:

Cumulative Distribution function:

Boxplot:
Continuous Random Variables

Probability Density function:

Cumulative Distribution function:

Boxplot:
Continuous Random Variables

Probability Density function:

$ f_X(x) $

Cumulative Distribution function:

$ F_X(x) = P(X \leq x) $

$ F_X(x) = \int_{-\infty}^{x} f_X(y) \, dy $

$ \frac{d}{dx} F_X(x) = f_X(x) $

Boxplot:
Continuous Random Variables

Probability Density function:

\[ f_X(x) \]

Cumulative Distribution function:

\[ F_X(x) = \int_{-\infty}^{x} f_X(y) \, dy \]

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Probability Density function:

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Continuous Random Variables

Probability Density function:

Cumulative Distribution function:

$$F_X(x) = P(X \leq x)$$

$$F_X(x) = \int_{-\infty}^{x} f_X(y) \, dy$$

$$\frac{d}{dx} F_X(x) = f_X(x)$$

Boxplot:
Continuous Random Variables

Probability Density function:

\[ f_X(x) \]

Cumulative Distribution function:

\[ F_X(x) = P(X \leq x) = \int_{-\infty}^{x} f_X(y) \, dy \]

\[ \frac{d}{dx} F_X(x) = f_X(x) \]

Boxplot:
Continuous Random Variables

Cumulative Distribution Function (CDF):

Discrete: \( F_X(x) = P(X \leq x) = \sum_{y=-\infty}^{x} p_X(y) \)

Continuous: \( F_X(x) = P(X \leq x) = \int_{-\infty}^{x} f_X(y) \, dy \)

Probability Density Function (PDF):

Discrete: \( p_X(x) = P(X = x) \quad \sum_{x=-\infty}^{\infty} p_X(x) = 1 \)

Continuous: \( f_X(x) = \frac{d}{dx} F_X(x) \quad \int_{-\infty}^{\infty} f_X(x) \, dx = 1 \)
Expectation and Variance

Motivational Example:

Experiment on Plant Growth (inbuilt R dataset)
- compares yields obtained under different conditions
Expectation and Variance

Motivational Example:

Experiment on Plant Growth (inbuilt R dataset)
- compares yields obtained under different conditions

- Compare means to test for differences
- Consider variance (and shape) of the distributions – help choose appropriate prior/protocol
- Assess uncertainty of parameter estimates – allow hypothesis testing
Expectation and Variance

Motivational Example:

Experiment on Plant Growth (inbuilt R dataset)
- compares yields obtained under different conditions

- Compare means to test for differences
- Consider variance (and shape) of the distributions – help choose appropriate prior/protocol
- Assess uncertainty of parameter estimates – allow hypothesis testing

In order to do any of this, we need to know how to describe distributions
i.e. we need to know how to work with descriptive statistics
Expectation and Variance

Discrete RV: \[ E(X) = \sum_{-\infty}^{\infty} xp_X(x) \]

Sample (empirical): \[ E(X) = \frac{1}{n} \sum_{i=1}^{n} x_i \] (explicit weighting not required)

Continuous RV: \[ E(X) = \int_{-\infty}^{\infty} xf_X(x)dx \]
Expectation and Variance

Normal Distribution:

\[ X \sim N(\mu, \sigma^2) \]

\[ f_X(x) = \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \]
Expectation and Variance

Normal Distribution:

\[ X \sim N(\mu, \sigma^2) \]

\[ f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \]

\[ E(X) = \int_{-\infty}^{\infty} xf_X(x) \, dx = \mu \]
Expectation and Variance

Standard Cauchy Distribution:
(also called Lorentz)

\[ f_X(x) = \frac{1}{\pi} \frac{1}{1 + x^2} \]

\[ E(X) = \int_{-\infty}^{\infty} xf_X(x) \, dx = \text{undefined} \]
Expectation and Variance

Expectation of a function of random variables:

\[
E(g(X)) = \int_{-\infty}^{\infty} g(x) f_X(x) dx
\]

Linearity:

\[
E(\alpha X + \beta) = \int_{-\infty}^{\infty} (\alpha x + \beta) f_X(x) dx
\]

\[
= \alpha \int_{-\infty}^{\infty} xf_X(x) dx + \beta \int_{-\infty}^{\infty} f_X(x) dx
\]

\[
= \alpha E(X) + \beta
\]
Expectation and Variance

Variance:

\[ \text{Var}(X) = E\left( (X - \mu)^2 \right) \]
\[ = E\left( (X - E(X))^2 \right) \]
\[ = E(X^2) - E(X)^2 \]

\[ X \sim N(0,1) \]
**Expectation and Variance**

Variance:

\[
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Expectation and Variance

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\[ X \sim N(0,1) \]
\[ X \sim N(0,2) \]
Expectation and Variance

Variance:

\[ \text{Var}(X) = E((X - \mu)^2) \]
\[ = E((X - E(X))^2) \]
\[ = E(X^2) - E(X)^2 \]

Population Variance:

\[ \sigma^2 = \frac{1}{n} \sum_{i=1}^{n} (x_i - \mu)^2 \]

Unbiased Sample Variance:

\[ s^2 = \frac{1}{n-1} \sum_{i=1}^{n} (x_i - \bar{x})^2 \]

\( X \sim N(0,1) \)
\( X \sim N(0,2) \)
Expectation and Variance

Variance:

\[ \text{Var}(X) = E(X^2) - E(X)^2 \]

Non-linearity:

\[ \text{Var}(\alpha X + \beta) = E((\alpha X + \beta)^2) - E(\alpha X + \beta)^2 \]

Standard deviation (s.d.): \( \sqrt{\text{Var}(X)} \)
Expectation and Variance

Variance:

\[ \text{Var}(X) = E(X^2) - E(X)^2 \]

Standard deviation (s.d.): \( \sqrt{\text{Var}(X)} \)

Non-linearity:

\[ \text{Var}(\alpha X + \beta) = E((\alpha X + \beta)^2) - E(\alpha X + \beta)^2 \]

\[ = E(\alpha^2 X^2 + 2\alpha \beta X + \beta^2) - (\alpha E(X) + \beta)^2 \]
Expectation and Variance

Variance:

\[ Var(X) = E(X^2) - E(X)^2 \]

Standard deviation (s.d.): \( \sqrt{Var(X)} \)

Non-linearity:

\[ Var(\alpha X + \beta) = E((\alpha X + \beta)^2) - E(\alpha X + \beta)^2 \]

\[ = E(\alpha^2 X^2 + 2\alpha \beta X + \beta^2) - (\alpha E(X) + \beta)^2 \]

\[ = (\alpha^2 E(X^2) + 2\alpha \beta E(X) + \beta^2) - (\alpha^2 E(X)^2 + 2\alpha \beta E(X) + \beta^2) \]
Expectation and Variance

Variance:

\[ \text{Var}(X) = E(X^2) - E(X)^2 \]

Standard deviation (s.d.): \( \sqrt{\text{Var}(X)} \)

Non-linearity:

\[
\text{Var}(\alpha X + \beta) = E\left((\alpha X + \beta)^2\right) - E(\alpha X + \beta)^2
\]
\[
= E\left(\alpha^2 X^2 + 2\alpha\beta X + \beta^2\right) - (\alpha E(X) + \beta)^2
\]
\[
= \left(\alpha^2 E(X^2) + 2\alpha\beta E(X) + \beta^2\right) - \left(\alpha^2 E(X)^2 + 2\alpha\beta E(X) + \beta^2\right)
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\[ \text{Var}(\alpha X + \beta) = E\left( (\alpha X + \beta)^2 \right) - E(\alpha X + \beta)^2 \]

\[ = E\left( \alpha^2 X^2 + 2\alpha \beta X + \beta^2 \right) - \left( \alpha E(X) + \beta \right)^2 \]

\[ = \left( \alpha^2 E(X^2) + 2\alpha \beta E(X) + \beta^2 \right) - \left( \alpha^2 E(X)^2 + 2\alpha \beta E(X) + \beta^2 \right) \]

\[ = \alpha^2 \left( E(X^2) - E(X)^2 \right) \]

\[ = \alpha^2 \text{Var}(X) \]
Expectation and Variance

Often data are standardised/normalised

Z-score/value: \[ Z = \frac{X - \mu}{\sigma} \]

Example:

\[ X \sim N(\mu, \sigma^2) \quad f_X(x) = \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \]

\[ Z \sim N(0,1) \quad f_Z(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \]
Moments

Shape descriptors
Moments

Shape descriptors

Li and Hartley (2006) Computer Vision
Saupe and Vranic (2001) Springer
Moments

Shape descriptors

Li and Hartley (2006) Computer Vision
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Moments

Moments provide a description of the shape of a distribution

<table>
<thead>
<tr>
<th>Raw moments</th>
<th>Central moments</th>
<th>Standardised moments</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\mu_1 = E(X) = \mu$ : mean</td>
<td>$0$</td>
<td>$0$</td>
</tr>
<tr>
<td>$\ldots$</td>
<td>$E((X - \mu)^2) = \sigma^2$ : variance</td>
<td>$1$</td>
</tr>
<tr>
<td>$\ldots$</td>
<td>$\ldots$</td>
<td>$E\left(\left(\frac{X-\mu}{\sigma}\right)^3\right)$ : skewness</td>
</tr>
<tr>
<td>$\ldots$</td>
<td>$\ldots$</td>
<td>$E\left(\left(\frac{X-\mu}{\sigma}\right)^4\right)$ : kurtosis</td>
</tr>
<tr>
<td>$\mu_n = E(X^n)$</td>
<td>$E((X - \mu)^n)$</td>
<td>$E\left(\left(\frac{X-\mu}{\sigma}\right)^n\right)$</td>
</tr>
</tbody>
</table>
Moments

Standard Normal:

Standard Log-Normal:
Moments

Moment generating function (MGF):

\[ M_X(t) = E(e^{Xt}) = 1 + tE(X) + \frac{t^2}{2!} E(X^2) + ... + \frac{t^n}{n!} E(X^n) + ... \]

Alternative representation of a probability distribution.

\[ \mu_n = E(X^n) = \frac{d^n}{dt^n} M_X(0) \]
Moments

Moment generating function (MGF):

\[ M_X(t) = E(e^{Xt}) = 1 + tE(X) + \frac{t^2}{2!} E(X^2) + \ldots + \frac{t^n}{n!} E(X^n) + \ldots \]

Alternative representation of a probability distribution.

\[ \mu_n = E(X^n) = \frac{d^n}{dt^n} M_X(0) \]

Example:

\[ X \sim N(\mu, \sigma^2) \quad \Rightarrow \quad M_X(t) = e^{t\mu + \frac{1}{2}\sigma^2 t^2} \]

\[ X \sim N(0,1) \quad \Rightarrow \quad M_X(t) = e^{\frac{1}{2}t^2} \]
Moments

However, MGF only exists if $E(X^n)$ exists

$$M_X(t) = E(e^{Xt})$$

Characteristic function always exists:

$$\varphi_X(t) = M_{iX}(t) = M_X(it) = E(e^{itX}) = \int_{-\infty}^{\infty} e^{itx} f_X(x) \, dx$$

Related to the probability density function via Fourier transform

Example: $X \sim N(0,1)$ \hspace{1cm} $\varphi_X(t) = e^{-t^2/2}$
The Law of Large Numbers (LLN)

Motivational Example:

Experiment on Plant Growth (inbuilt R dataset)
  - compare yields obtained under different conditions

- Want to estimate the population mean using the sample mean.

- How can we be sure that the sample mean reliably estimates the population mean?
The Law of Large Numbers (LLN)

Does the sample mean reliably estimate the population mean?

The Law of Large Numbers:

\[ \bar{X}_n = \frac{1}{n} \sum_{i=1}^{n} X_i \xrightarrow{n \to \infty} \mu \]

Providing \( X_i : \text{i.i.d.} \)
The Law of Large Numbers (LLN)

Does the sample mean reliably estimate the population mean?

The Law of Large Numbers:

\[
\bar{X}_n = \frac{1}{n} \sum_{i=1}^{n} X_i \xrightarrow{n \to \infty} \mu
\]

Providing \(X_i : \text{i.i.d.}\)

\(X \sim U(0,1)\)

\(\mu = 0.5\)
The Central Limit Theorem (CLT)

Question – given a particular sample, thus known sample mean, how reliable is the sample mean as an estimator of the population mean?

Furthermore, how much will getting more data improve the estimate of the population mean?

Related question – given that we want the estimate of the mean to have a certain degree of reliability (i.e. sufficiently low S.E.), how many observations do we need to collect?

The Central Limit Theorem helps answer these questions by looking at the distribution of stochastic fluctuations about the mean as $n \to \infty$
The Central Limit Theorem (CLT)

The Central Limit Theorem states:

For large $n$: \[ \sqrt{n}(\bar{X}_n - \mu) \sim N(0, \sigma^2) \]

Or equivalently:
\[ \bar{X}_n \sim N(\mu, \frac{\sigma^2}{n}) \]

More formally:
\[ \sqrt{n}\left(\frac{1}{n} \sum_{i=1}^{n} X_i - \mu\right) \xrightarrow{d} N(0, \sigma^2) \]

Conditions:
\[ E(X_i) = \mu \]
\[ Var(X_i) = \sigma^2 < \infty \]
\[ X_i : \text{i.i.d. RVs (any distribution)} \]
The Central Limit Theorem (CLT)

The Central Limit Theorem states:

For large \( n \):

\[
\sqrt{n}(\bar{X}_n - \mu) \sim N(0, \sigma^2)
\]

\[
\bar{X}_n \sim N\left(\mu, \frac{\sigma^2}{n}\right)
\]

\( X \sim U(0, 1) \quad \mu = 0.5 \quad \sigma^2 = \frac{1}{12} \)
The Central Limit Theorem (CLT)

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The Central Limit Theorem (CLT)

The Central Limit Theorem states:

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Proof of the Central Limit Theorem:

$$\sqrt{n}\left(\frac{1}{n} \sum_{i=1}^{n} X_i - \mu\right) \xrightarrow{d} N(0, \sigma^2)$$
The Central Limit Theorem (CLT)

Proof of the Central Limit Theorem:

\[ \sqrt{n} \left( \frac{1}{n} \sum_{i=1}^{n} X_i - \mu \right) \xrightarrow{d} N(0, \sigma^2) \]

\[ \frac{1}{\sqrt{n}} \left( \sum_{i=1}^{n} X_i - n\mu \right) \xrightarrow{d} N(0, \sigma^2) \]
The Central Limit Theorem (CLT)

Proof of the Central Limit Theorem:

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\[ \frac{1}{\sigma \sqrt{n}} \left( \sum_{i=1}^{n} X_i - n \mu \right) \xrightarrow{d} N(0, 1) \]
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Proof of the Central Limit Theorem:

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\[
\frac{1}{\sqrt{n}} \sum_{i=1}^{n} \left( \frac{X_i - \mu}{\sigma} \right) \xrightarrow{d} N(0, 1)
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\[ \frac{1}{\sqrt{n}} \sum_{i=1}^{n} \left( \frac{X_i - \mu}{\sigma} \right) \xrightarrow{d} N(0, 1) \]

\[ \sum_{i=1}^{n} \frac{Z_i}{\sqrt{n}} \xrightarrow{d} N(0, 1) \]

\[ Z_i = \frac{X_i - \mu}{\sigma} \]
The Central Limit Theorem (CLT)

Proof of the Central Limit Theorem:

$$\sum_{i=1}^{n} \frac{Z_i}{\sqrt{n}} \xrightarrow{d} N(0,1)$$
The Central Limit Theorem (CLT)

Proof of the Central Limit Theorem:

\[ \sum_{i=1}^{n} \frac{Z_i}{\sqrt{n}} \xrightarrow{d} N(0,1) \]

\[ \varphi_n(t) = \varphi \left( \frac{t}{\sqrt{n}} \right) \]

\[ E(e^{i \frac{Z}{\sqrt{n}} t}) = E(e^{iZ \frac{t}{\sqrt{n}}}) \]
The Central Limit Theorem (CLT)

Proof of the Central Limit Theorem:

$$\sum_{i=1}^{n} \frac{Z_i}{\sqrt{n}} \xrightarrow{d} N(0,1)$$

$$\varphi_n(t) = \varphi_n \left( \frac{t}{\sqrt{n}} \right)$$

$$= \prod_{i=1}^{n} \varphi_{Z_i} \left( \frac{t}{\sqrt{n}} \right)$$

$$E(e^{it\left(\frac{Z}{\sqrt{n}}\right)}) = E(e^{it\left(\frac{Z}{\sqrt{n}}\right)})$$

$$E(e^{it(Z_1+Z_2)}) = E(e^{itZ_1})E(e^{itZ_2})$$
The Central Limit Theorem (CLT)

Proof of the Central Limit Theorem:

\[
\sum_{i=1}^{n} \frac{Z_i}{\sqrt{n}} \xrightarrow{d} N(0, 1)
\]

\[
\varphi_n(t) = \varphi \left( \sum_{i=1}^{n} \frac{Z_i}{\sqrt{n}} \right) = \prod_{i=1}^{n} \varphi_{Z_i} \left( \frac{t}{\sqrt{n}} \right) = \left( \varphi \left( \frac{t}{\sqrt{n}} \right) \right)^n
\]

\[
E(e^{i\frac{Z}{\sqrt{n}}t}) = E(e^{iz\left( \frac{t}{\sqrt{n}} \right)})
\]

\[
E(e^{it(Z_1+Z_2)}) = E(e^{itZ_1})E(e^{itZ_2})
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The Central Limit Theorem (CLT)

Proof of the Central Limit Theorem:

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\sum_{i=1}^{n} \frac{Z_i}{\sqrt{n}} \xrightarrow{d} N(0,1)
\]

\[
\varphi^n \left( \sum_{i=1}^{n} \frac{Z_i}{\sqrt{n}} \right) = \varphi^n \left( \sum_{i=1}^{n} \frac{t}{\sqrt{n}} \right)
\]

\[
= \prod_{i=1}^{n} \varphi_{Z_i} \left( \frac{t}{\sqrt{n}} \right)
\]

\[
= \left( \varphi_{Z} \left( \frac{t}{\sqrt{n}} \right) \right)^n = \left( 1 - \frac{t^2}{2n} + o \left( \frac{t^2}{n} \right) \right)^n
\]

\[
E(e^{i \left( \frac{Z}{\sqrt{n}} \right) t}) = E(e^{iz \left( \frac{t}{\sqrt{n}} \right)})
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\[
E(e^{it(Z_1 + Z_2)}) = E(e^{itZ_1})E(e^{itZ_2})
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The Central Limit Theorem (CLT)

Proof of the Central Limit Theorem:

$$\sum_{i=1}^{n} \frac{Z_i}{\sqrt{n}} \xrightarrow{d} N(0,1)$$

$$\varphi_n(t) = \varphi \left( \frac{t}{\sqrt{n}} \right)$$

$$= \prod_{i=1}^{n} \varphi_{Z_i} \left( \frac{t}{\sqrt{n}} \right)$$

$$= \left( \varphi \left( \frac{t}{\sqrt{n}} \right) \right)^n$$

$$= \left( 1 - \frac{t^2}{2n} + o \left( \frac{t^2}{n} \right) \right)^n \rightarrow e^{-t^2/2}$$

$$E(e^{itZ}) = E(e^{i(Z + Z)}) = E(e^{itZ_1})E(e^{itZ_2})$$

$$e^x = \lim_{n \to \infty} \left( 1 + \frac{x}{n} \right)^n$$
Summary

Considered how:

• Probability Density Functions (PDFs) and Cumulative Distribution Functions (CDFs) are related, and how they differ in the discrete and continuous cases

• Expectation is at the core of Statistical theory, and Moments can be used to describe distributions

• The Central Limit Theorem identifies how/why the Normal distribution is fundamental

The Normal distribution is also popular for other reasons:

• Maximum entropy distribution (given mean and variance)

• Intrinsically related to other distributions ($t$, $F$, $\chi^2$, Cauchy, …)

• Also, it is easy to work with
Countless books + online resources!

Probability and Statistical theory:


General comprehensive introduction to (almost) everything mathematics:

• Garrity (2002) All the mathematics you missed: but need to know for graduate school. Cambridge University Press.